Uncertainty Quantification in Machine Learning From Aleatoric to Epistemic

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Need for uncertainty-awareness of ML systems



Lack of uncertainty-awareness of ML systems

Predictions by EfficientNet on test images from ImageNet: For the left image, the neural network predicts "typewriter keyboard" with certainty 83.14%, for the right image "stone wall" with certainty 87.63%.





Uncertainty representation and levels of uncertainty-awareness



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Aleatoric versus epistemic uncertainty

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 - refers to the notion of randomness, that is, the variability in the outcome which is due to inherently random effects,
 - is a property of the data-generating process,
 - and as such irreducible.

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 - is a property of the data-generating process,
 - and as such irreducible.
- **Epistemic** (aka systematic) uncertainty
 - refers to uncertainty caused by a lack of knowledge, i.e.,
 - to the epistemic state of the **agent** (e.g., learning algorithm),
 - ▶ can in principle be reduced on the basis of additional information (e.g., training data).

Aleatoric versus epistemic uncertainty



"Not knowing the chance of mutually exclusive events and knowing the chance to be equal are two quite different states of knowledge"

Ronald Fisher (1890-1962)



Aleatoric versus epistemic uncertainty in ML

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■ ... but also on the underlying model assumptions:



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- Predict the next number: 116, 304, 194, 341, 224, 654, 609, 625, 533, 91, 205, 35, 527, 611, 128, 235, 348, 912, 582, 52, 672, 20, 856, 904, 628, 273, 615, 105, 610, 862, 384, 705, 73, 794, 775, 156, ??

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 $x \leftarrow x imes 237 \mod 971$







Sources of uncertainty



 h^*

Agenda

- 1. Aleatoric and epistemic uncertainty
- 2. Learning uncertainty-aware predictors
- 3. Uncertainty quantification
- 4. Summary and outlook

Predictive uncertainty

Predictive uncertainty

We assume a standard setting of supervised learning and are mainly interested in (per-instance) predictive uncertainty, i.e., the uncertainty in a prediction

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■ Various approaches have been proposed in the literature:

- Capture model uncertainty, translate into predictive uncertainty
- Validation and self-assessment
- Direct uncertainty prediction

The Bayesian approach

- A Bayesian learner maintains a probability distribution over the hypothesis space.
- The less concentrated that distribution, the higher the learner's epistemic uncertainty.



Posterior predictive distribution



Ensemble methods



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- Example: Estimation of error rate via (cross-)validation.
- Yet, this is a **global** performance measure, not **per-instance**.
- Per-instance uncertainty estimation appears to be difficult and indeed has theoretical limits (Barber *et al.*, 2021).





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- On calibration data, the learner extracts information such as: A predicted probability of ≈ 0.6 actually means a true probability of ≈ 0.3 .
- **Grouping** of instances with same score (predicted probability), needed to construct frequentist corrections of level-1 predictions based on level-0 data.
- A calibrator is a **one-dimensional** function, hence easier to learn.

Conformal prediction



- A conformal predictor uses calibration data to learn rules such as: With high probability, true outcomes have a nonconformity of at most α₀.
- This allows for constructing non-trivial yet valid **prediction sets**.
Level-2 predictions



Previous approaches refer to level-1 uncertainty, though level-2 estimation is in principle also possible (e.g., Venn predictors)

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- Yet, learning such a predictor appears to be difficult (and also includes learning of f*(x) or knowledge thereof).
- Besides, one may question the definition of **uncertainty** in terms of **loss**.

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Direct prediction



Given training data $\mathcal{D} = \{(\mathbf{x}_i, y_i)\}_{i=1}^N \subset \mathcal{X} \times \mathcal{Y}$, can we train a predictor

 $\hat{h}: \mathcal{X} \longrightarrow \mathbb{P}ig(\mathbb{P}(\mathcal{Y})ig)$

via (variants of) empirical risk minimisation (ERM), i.e.,

$$\hat{h} = \operatorname*{arg\,min}_{h} \sum_{i=1}^{N} \ell_2\left(\hat{h}(\boldsymbol{x}_i), y_i\right) ,$$

with a suitable level-2 loss function

$$\ell_2: \mathbb{P}(\mathbb{P}(\mathcal{Y})) \times \mathcal{Y} \longrightarrow \mathbb{R},$$

such that the predictor represents its epistemic uncertainty in a faithful way?

Example: predicting a Dirichlet distribution



The case of level-1 predictions



 \boldsymbol{x}

Training a probabilistic predictor via empirical risk minimisation, i.e.,

$$\hat{h} = \operatorname*{arg\,min}_{h} \sum_{i=1}^{N} \ell_1\left(\hat{h}(\boldsymbol{x}_i), y_i\right) \,,$$

yields good (unbiased) predictors if ℓ_1 is a (strictly) **proper scoring rule**, which incentivises the learner to predict the true p(y | x).

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■ A loss function $\ell_1 : \mathbb{P}(\mathcal{Y}) \times \mathcal{Y} \longrightarrow \mathbb{R}$ is a proper scoring rule if the expected loss minimiser coincides with the true probability **p**:

$$oldsymbol{
ho} = rgmin_{oldsymbol{\hat{p}}} \mathbb{E}_{Y \sim oldsymbol{
ho}} \, \ell_1(oldsymbol{\hat{p}}, Y)$$

A scoring rule is **strictly proper** if the minimiser is unique.

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$$\ell_2(Q, y) = \mathbb{E}_{P \sim Q} \,\ell_1(P, y) \;,$$

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- Examples of level-1 losses include cross entropy (Charpentier *et al.*, 2020) and Brier score (Sensoy *et al.*, 2018).
- Besides, a regularised version has been proposed:

$$\ell_{2}(Q, y) = \mathbb{E}_{P \sim Q} \ell_{1}(P, y) + \lambda d_{KL}(Q, Q_{0})$$

Appropriate level-2 losses

Informally, we define a level-2 loss function ℓ_2 as **appropriate** if the following holds for the empirical loss minimiser

$$Q^{(N)} = rgmin_Q rac{1}{N} \sum_{n=1}^N \ell_2\left(Q, y^{(n)}
ight).$$

on any i.i.d. observational data sequence $y^{(1)}, y^{(2)}, \ldots$ with $y^{(i)} \sim P^*$:

- (A1) The learner's **uncertainty gradually decreases** (in expectation) with increasing sample size N, in terms of a suitable uncertainty measure U.
- (A2) In the limit $N \to \infty$, all epistemic uncertainty disappears and $Q^{(N)} \to \delta_{P^*}$.

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- The results are general in the sense that Q can be any level-2 distribution, not necessarily restricted to Dirichlet distributions.
- Moreover, the results do not depend on the underlying uncertainty measure U, as long as U is not constant, maximal for the uniform distribution and minimal for Dirac measures.

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- The results are general in the sense that Q can be any level-2 distribution, not necessarily restricted to Dirichlet distributions.
- Moreover, the results do not depend on the underlying uncertainty measure U, as long as U is not constant, maximal for the uniform distribution and minimal for Dirac measures.
- The results reveal that the quality of a (level-2) prediction Q cannot be judged solely in the context of (level-0) observations y.

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■ Given a prediction *h*(*x*) in the form of a second-order distribution or a credal set, how to quantify the **total uncertainty** in that prediction in terms of a single number?



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• We may also seek a **decomposition** into an aleatoric and an epistemic part:

TU = AU + EU

• One idea is to quantify the different types of uncertainty in terms of

Shannon entropy

$$H[Y] = -\sum_{y \in \mathcal{Y}} \mathbf{p}(y) \log_2 \mathbf{p}(y),$$

conditional entropy

$$H[Y | P] = -\int Q(p | D) \left(\sum_{y \in \mathcal{Y}} \mathbf{p}(y | p) \log_2 \mathbf{p}(y | p) \right) dp,$$

and mutual information

between outcome Y and (level-1) distribution P (Malinin and Gales, 2018), respectively:



Remarks

- MI actually measures the (average) divergence between the candidate (level-1) distributions, so it is rather a measure of conflict than ignorance (which is difficult to capture in terms of probabilities anyway).
- One may also question the additive decomposition TU = AU + EU itself.



Uncertainty of credal sets

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■ Uncertainty measures *U* for credal sets have been studied axiomatically:

- Al **Non-negativity, range**: U is non-negative and upper-bounded by some value $r \in \mathbb{R}$, for example $r = \log(K)$, which is assumed for $Q = \Delta_K$ (the case of complete ignorance).
- A2 **Continuity**: U is a continuous functional.
- A3 Monotonicity: If $Q \subseteq Q'$ for credal sets Q, Q', then $U(Q) \leq U(Q')$.
- A4 **Probability consistency**: U reduces to standard Shannon entropy in the case where Q reduces to a single probability distribution.
- A5 **Sub-additivity**: For a (joint) credal set Q on a product space $\mathcal{Y}' \times \mathcal{Y}''$ with marginals Q' resp. Q'',

$$U(Q) \leq U(Q') + U(Q'')$$
.

A6 Additivity: The last inequality is an equality in the case where Q' and Q'' are independent (assuming a suitably defined notion of independence).

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A well-founded measure of epistemic uncertainty is the generalised Hartley measure

$$\mathsf{GH}(Q) := \sum_{A \subseteq \mathcal{Y}} \mathsf{m}_Q(A) \, \mathsf{log}(|A|) \,,$$

which extends the Hartley measure $H(A) := \log(|A|)$ from sets to graded sets.

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Although an equally well-justified measure of aleatoric uncertainty (conflict) in the form of an extension of Shannon entropy has not been found so far (Klir, 2005), the lower entropy is a natural measure of irreducible uncertainty:

$$S_*(Q) := \min_{q \in Q} S(q)$$

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■ Idea: Fix two "good" measures and derive the third one in terms of the difference.

$$S^*(Q) = \left(\underbrace{S^*(Q) - GH(Q)}_{GS(Q)}\right) + GH(Q)$$
$$S^*(Q) = S_*(Q) + \left(S^*(Q) - S_*(Q)\right)$$

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H. et al. (2022) provide a critical discussion of such decompositions and isolate potential deficiencies.

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- We proposed and axiomatically justified a new measure of total predictive uncertainty, more tailored to the ML setting, as well as its decomposition into aleatoric and epistemic uncertainty.
- In the case of binary classification, where a credal prediction is of the form

$$Q_{\alpha,\beta} = \left\{ \operatorname{Bern}(p) \, | \, \alpha \leq p \leq \beta \right\},$$

the measure is given as follows:

_

$$\mathsf{\Gamma}\mathsf{P}(\alpha,\beta) = \underbrace{\mathsf{min}(1-\alpha,\beta)}_{\mathsf{total}} = \underbrace{\mathsf{min}(\alpha,1-\beta)}_{\mathsf{aleatoric}} + \underbrace{(\beta-\alpha)}_{\mathsf{epistemic}}$$

Empirical evaluation

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Ensemble-based construction of credal predictions.

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- **Ensemble-based construction** of credal predictions.
- Accuracy-rejection curves: Allow the learner to reject the *r*% presumably most uncertain test cases and measure accuracy on the remaining ones.



Results

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- Empirical results match with theory: Formally justified measures show strong performance, whereas the "derived" measures perform very poorly.
- Newly proposed measure yields the only decomposition of total into aleatoric and epistemic uncertainty, such that all three measures produce meaningful results.



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- Quantifying predictive uncertainty in a theoretically sound manner, and disentangling total into aleatoric and epistemic uncertainty, is difficult, too.

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- Quantifying predictive uncertainty in a theoretically sound manner, and disentangling total into aleatoric and epistemic uncertainty, is difficult, too.
- Usefulness of generalized uncertainty calculi?

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Model misspecification



What uncertainty should the learner report at x?

Example: level-2 distributions over Bernoulli







TU = 1.00, AU = 0.93, EU = 0.07



TU = 0.76, AU = 0.69, EU = 0.07



TU = 1.00, AU = 0.96, EU = 0.04



TU = 0.73, AU = 0.67, EU = 0.07



TU = 1.00, AU = 0.00, EU = 1.00

Evaluation: accuracy-rejection curves

Reject test instances for which (total, aleatoric, epistemic) uncertainty exceeds a certain threshold, measure accuracy on the remaining ones.



Theorem 1. If $\ell_1 : \mathbb{P}(\mathcal{Y}) \times \mathcal{Y} \longrightarrow \mathbb{R}$ is such that

 $\ell_1\left(\mathbb{E}_{\boldsymbol{\theta}\sim Q}[\boldsymbol{\theta}], y\right) \leq \mathbb{E}_{\boldsymbol{\theta}\sim Q}\left[\ell_1\left(\boldsymbol{\theta}, y\right)\right]$

for all $y \in \mathcal{Y}$, then $\ell_2(Q, y) = \mathbb{E}_{\theta \sim Q} \left[\ell_1(\theta, y) \right]$ violates A1.

- Condition on ℓ_1 is fulfilled if ℓ_1 is convex (in the first argument)
- Includes Brier score and cross-entropy, which are (strictly) convex
- Proof reveals that \hat{Q} is always a point-mass on $\mathbb{P}(\mathcal{Y})$ (i.e., a level-1 prediction)

Theorem 2. If $\ell_1 : \mathbb{P}(\mathcal{Y}) \times \mathcal{Y} \longrightarrow \mathbb{R}$ is strictly convex in its first argument, then there exists $\overline{\lambda} > 0$ such that

 $\ell_2(Q, y) = \mathbb{E}_{\theta \sim Q} \left[\ell_1(\theta, y) \right] + \lambda \cdot \mathrm{KL} \left[Q, \mathrm{Unif}(\mathbb{P}(\mathcal{Y})) \right]$

violates A1 for all $\lambda < \overline{\lambda}$.

Theorem 3. If $\ell_1 : \mathbb{P}(\mathcal{Y}) \times \mathcal{Y} \longrightarrow \mathbb{R}$ is

(i) strictly proper,

(ii) locally Lipschitz (in the first argument),

then there exists $\underline{\lambda} > 0$ such that

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Brier score and cross-entropy fulfill both (i) and (ii).

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Here, one might be quite sure about the class of the query under standard assumptions of binary classification, but much less so in a setting of **novelty detection**, where new classes may emerge.

The Dirichlet distribution

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- **\blacksquare** The probability density function is defined on the K simplex

$$\Delta_{\mathcal{K}} = \left\{ \boldsymbol{\theta} = (\theta_1, \dots, \theta_{\mathcal{K}})^\top \, | \, \theta_1, \dots, \theta_{\mathcal{K}} \geq 0, \, \sum_{k=1}^{\mathcal{K}} \theta_k = 1 \right\}$$

and given as follows:

$$p(\boldsymbol{ heta} \mid \boldsymbol{lpha}) = p(heta_1, \dots, heta_K \mid \boldsymbol{lpha}) = rac{1}{\mathbb{B}(\boldsymbol{lpha})} \prod_{k=1}^K heta_k^{lpha_k - 1},$$

where the normalisation constant is the multivariate Beta function.

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where the normalisation constant is the multivariate Beta function.

In Bayesian statistics, the Dirichlet distribution is commonly used as the conjugate prior of the multinomial distribution.