



POLITECNICO  
MILANO 1863



# Tailoring the Tails: Enhancing the Reliability of Probabilistic Load Forecasts

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a joint work with **Roberto Baviera**, Politecnico di Milano

**COPA 2024 Conference**

Milan, 10 September 2024

# Probabilistic Load Forecasting

The time series displays **seasonality on multiple time scales**

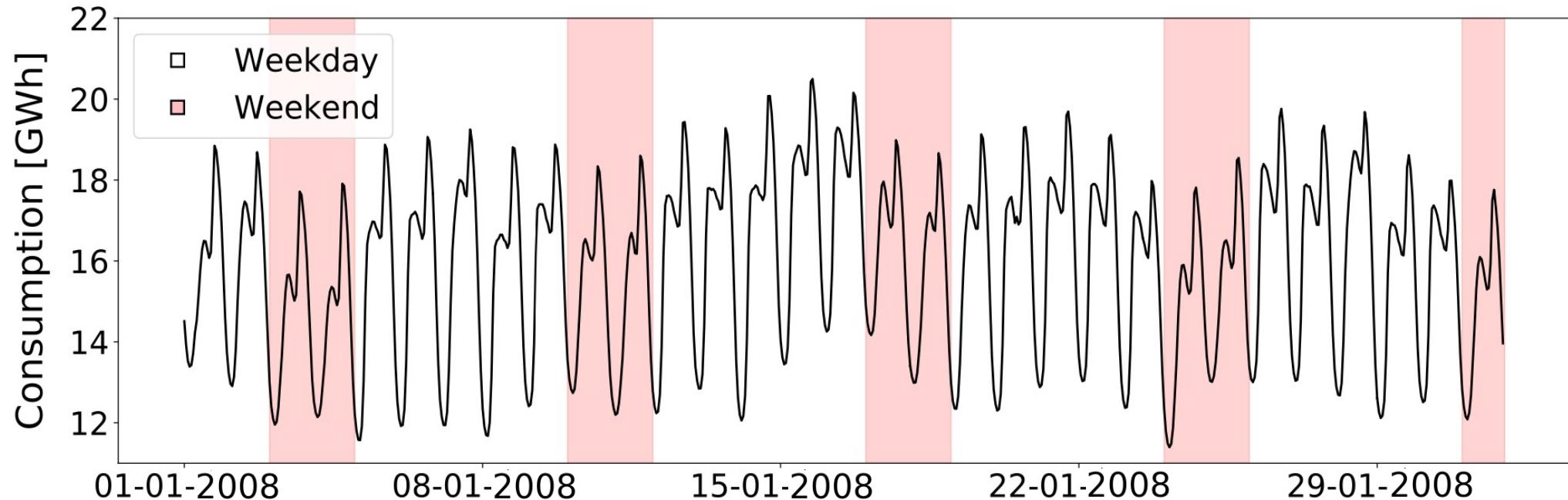


Figure: **Hourly** aggregate consumption of New England

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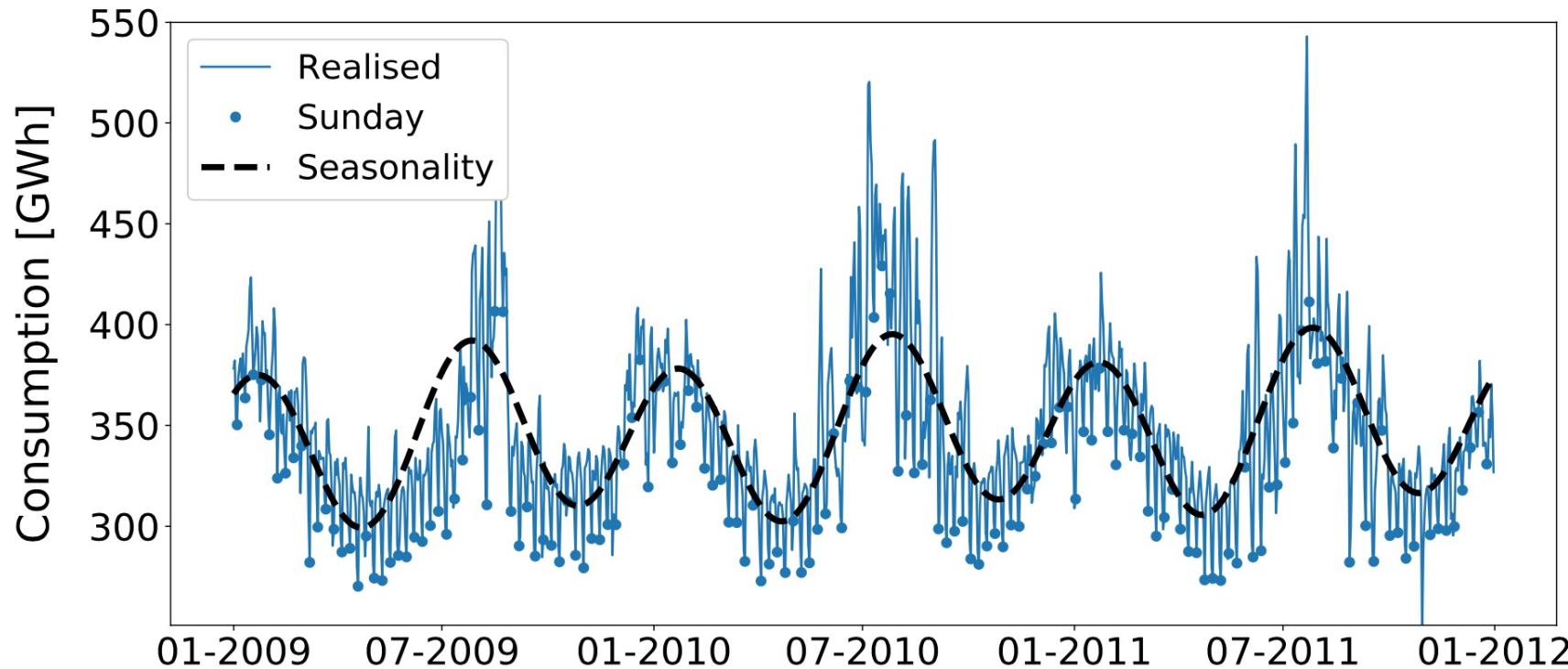


Figure: **Daily** aggregate consumption of New England

# Probabilistic Load Forecasting

The time series displays evident **seasonal heteroskedasticity**

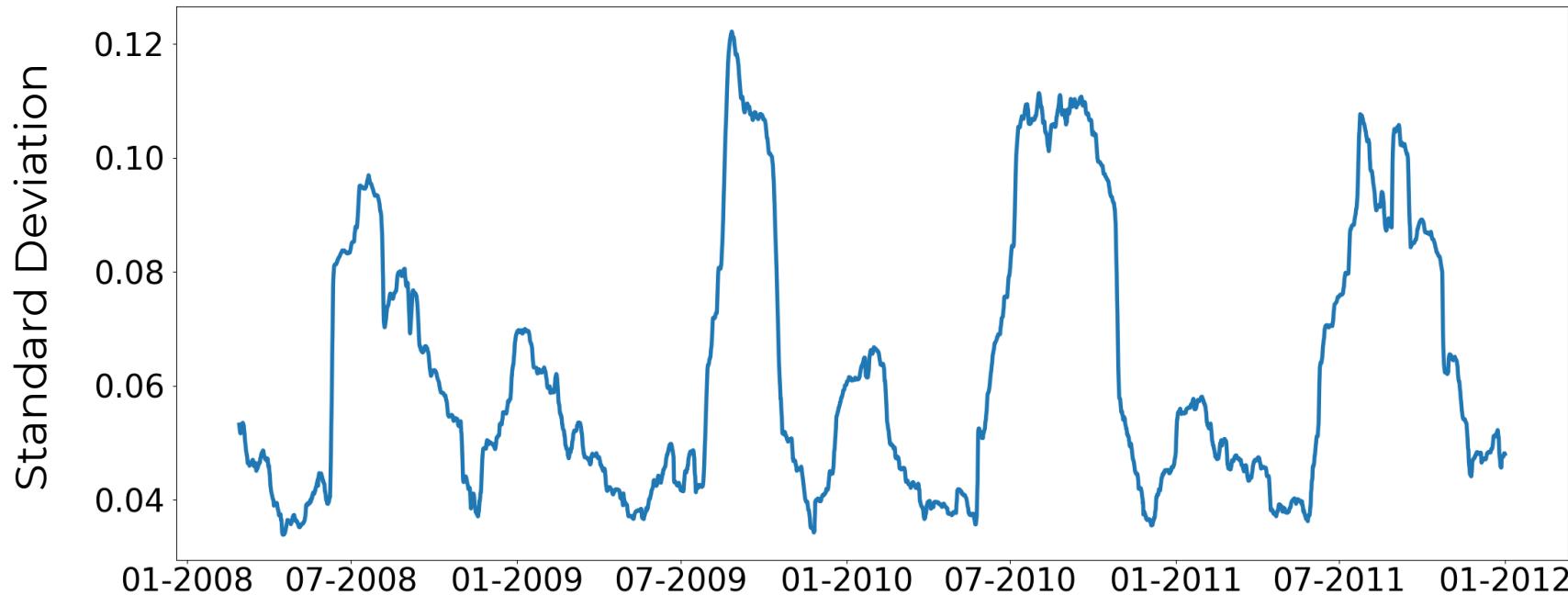


Figure: **Standard Deviation** of deseasonalised time series of consumption (on a 2-months rolling window)

# Probabilistic Load Forecasting



- It plays a vital role in decision-making processes: unit commitment, reserve management, economic dispatch and maintenance scheduling
- Forecast errors may result in grid instability, significant financial losses and potential blackouts.

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It is fundamental to take into account predictive **overconfidence**



overestimate **likely** events



underestimate **unlikely** events

# Research questions



How can we generate probabilistic forecasts  
for highly **non**-exchangeable time series data?



How can we manage **overconfidence** in  
probabilistic load forecasting?

# What does **overconfidence** look like?

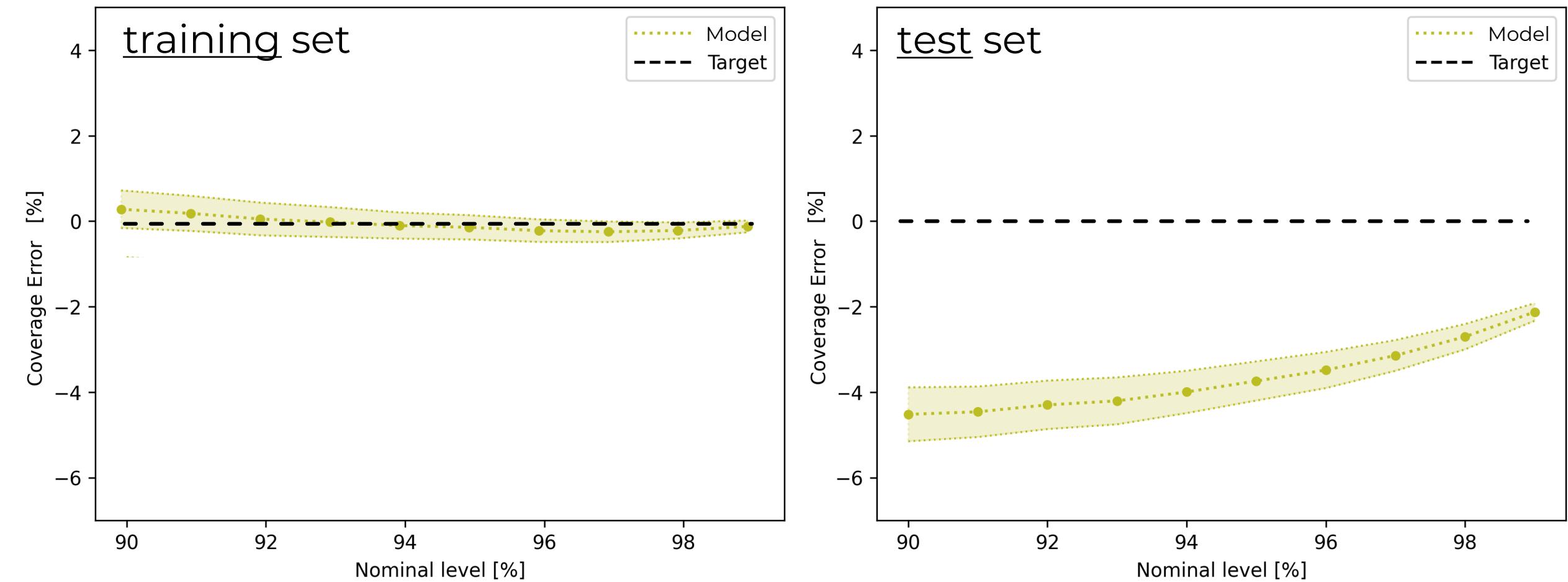


Figure: Prediction Intervals coverage plots on training and test sets  
Bands are intended to represent  $SE$  for 10 repetition with different seed

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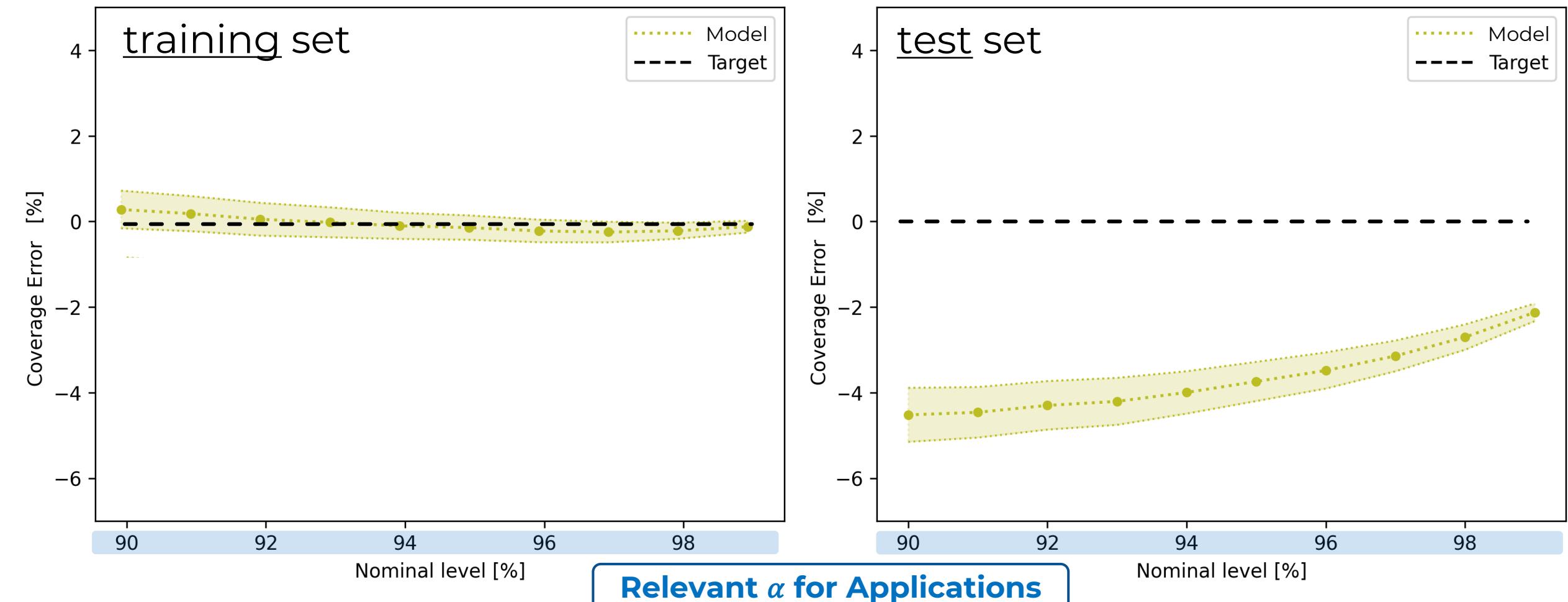


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# Outline of the talk

- ✓ Introduction and motivation

- The proposed methodology

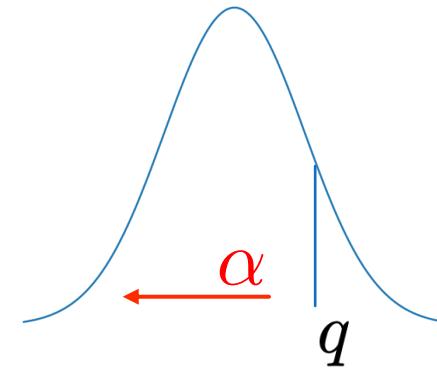
- Results

- Conclusions

# Pinball Loss

- Pinball Loss (PL; Konker & Bassett, 1978) – Evaluating Single Quantiles

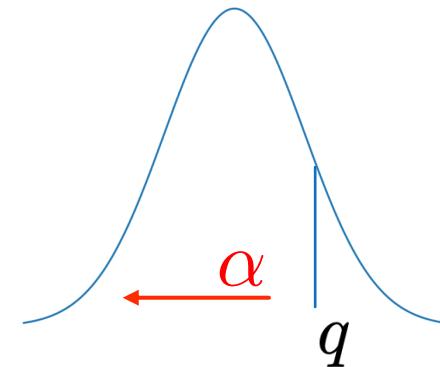
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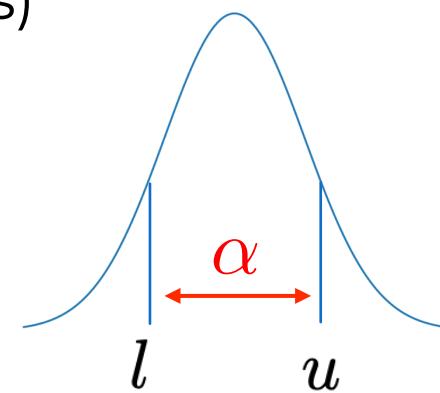
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- Central Pinball Loss – Adapting PL for Prediction Intervals (PIs)

$$\mathcal{P}_C(l, u, y; \alpha) := \mathcal{P}\left(l, y; \frac{1 - \alpha}{2}\right) + \mathcal{P}\left(u, y; \frac{1 + \alpha}{2}\right)$$

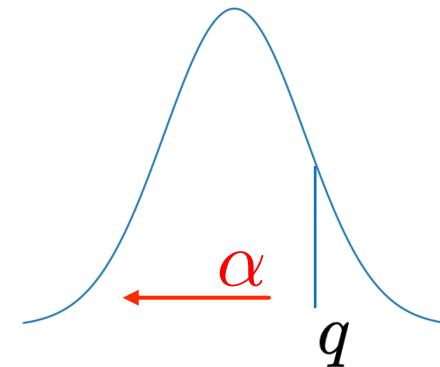


$$\frac{1 - \alpha}{2} \quad \frac{1 + \alpha}{2}$$

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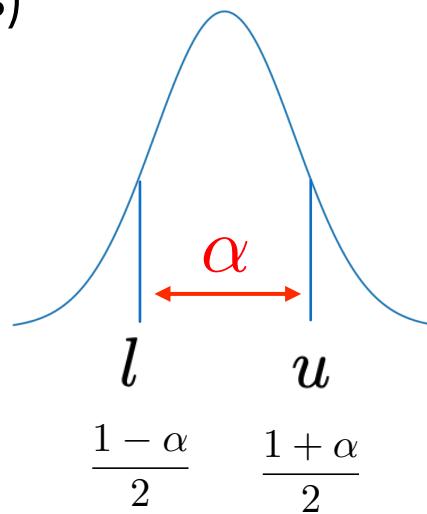
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# From PI to Distributions: CRPS

- When a probabilistic forecasting model issues a CDF  $F$ 
  - average Central Pinball Loss over all confidence levels  
(see, e.g., Gneiting & Raftery, 2007)

**Continuous Ranked Probabilistic Score** (Matheson & Winkler, 1986)

$$\text{CRPS}(F, y) = \int_0^1 \mathcal{P}_C \left( F^{-1} \left( \frac{1-\alpha}{2} \right), F^{-1} \left( \frac{1+\alpha}{2} \right), y ; \alpha \right) d\alpha$$

  
**Extrema of the  $\alpha$ -PI**

- Often used for forecast assessment, less often as a loss function

# Tackling predictive **overconfidence**

**The goal:** building a new robust loss function to be used at training time

$$\mathcal{P}_C(l, u, y; \alpha) = \frac{1 - \alpha}{2}(u - l) + \begin{cases} (y - u) & \text{if } y > u, \\ 0 & \text{if } y \in [l, u], \\ (l - y) & \text{if } y < l. \end{cases}$$

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- For any  $\lambda \in [0, 1)$ , we define the  **$\lambda$ -adjusted** Central Pinball Loss as

$$\mathcal{P}_C^{[\lambda]}(l, u, y; \alpha) = (1 - \lambda) \frac{1 - \alpha}{2}(u - l) + \begin{cases} (y - u) & \text{if } y > u, \\ 0 & \text{if } y \in [l, u], \\ (l - y) & \text{if } y < l. \end{cases}$$

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Computing the integral average of the new loss measure over quantiles:

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- This scoring rule is well-defined under mild integrability assumptions for the CDF and can be explicitly computed for all relevant distributions.

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**Proposition:** Let  $F$  be the CDF of a Gaussian with mean  $\mu$  and variance  $\sigma^2$ . Then:

$$\mathcal{G}\text{-CRPS}^{[\lambda]}(\mu, \sigma, y) := \sigma \left[ \frac{y - \mu}{\sigma} \left( 2\mathcal{N} \left( \frac{y - \mu}{\sigma} \right) - 1 \right) + 2\varphi \left( \frac{y - \mu}{\sigma} \right) - \frac{\lambda + \sqrt{2}(1 - \lambda)}{\sqrt{\pi}} \right]$$

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**CORRECTION TERM**

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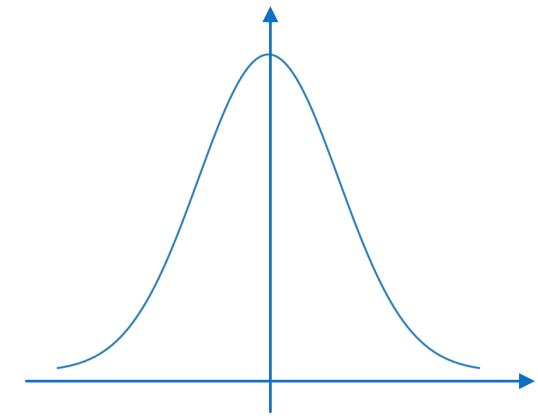
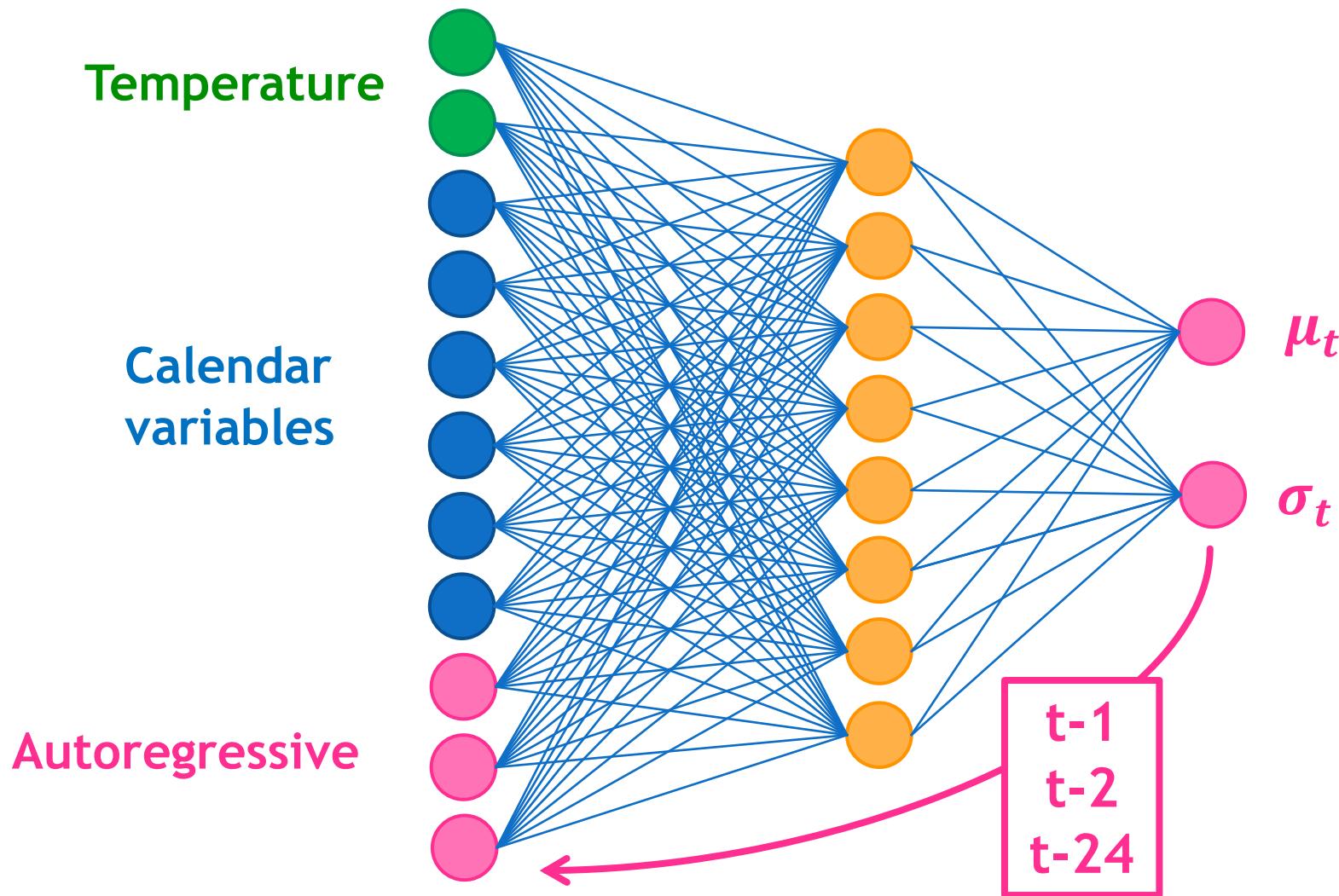
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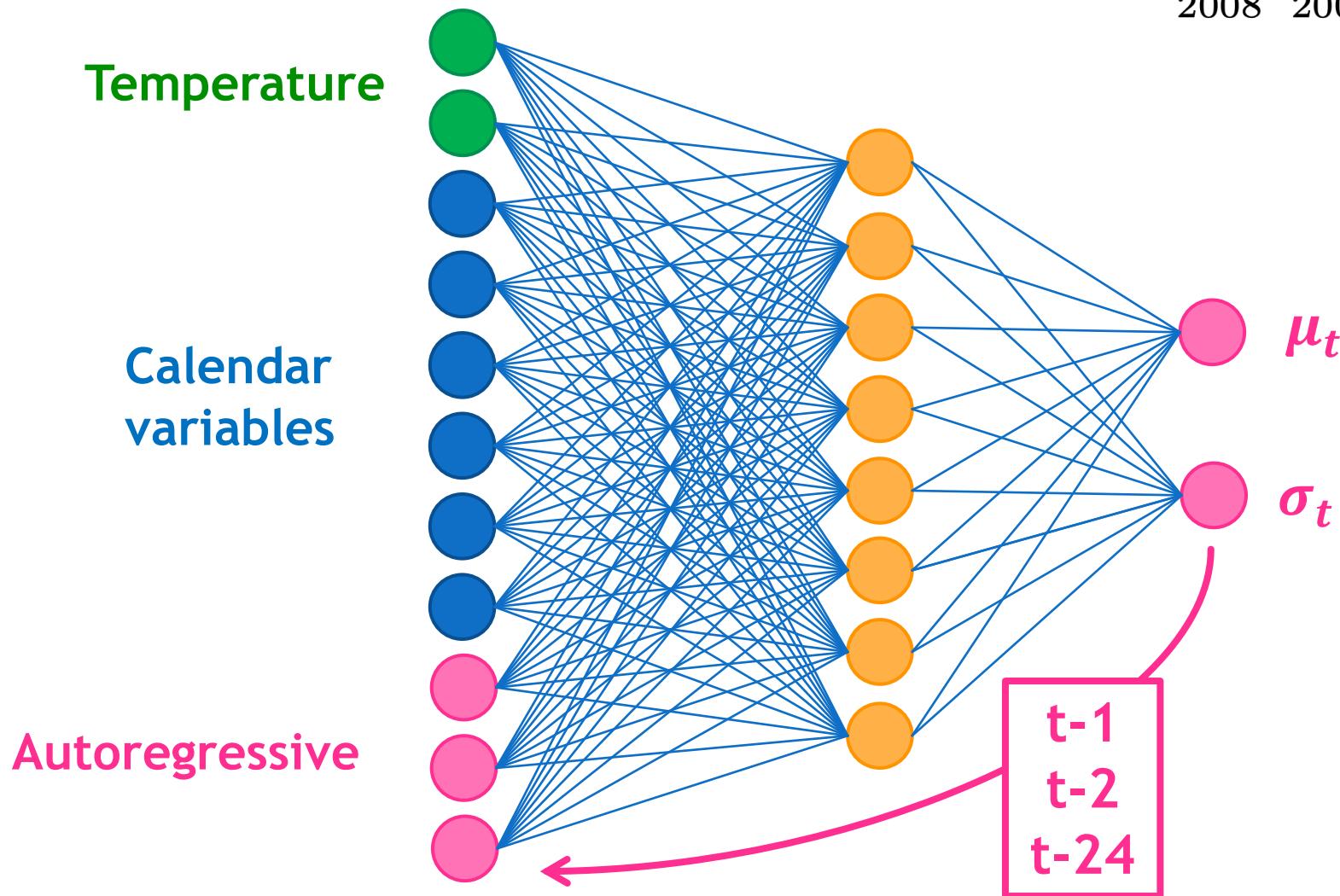
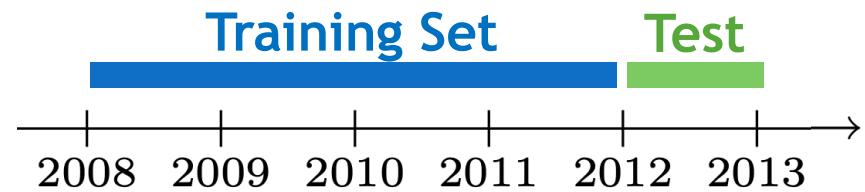
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# Modelling Approach



[Azzone & Baviera (2021),  
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# PIs Backtesting (on test set, 2012)

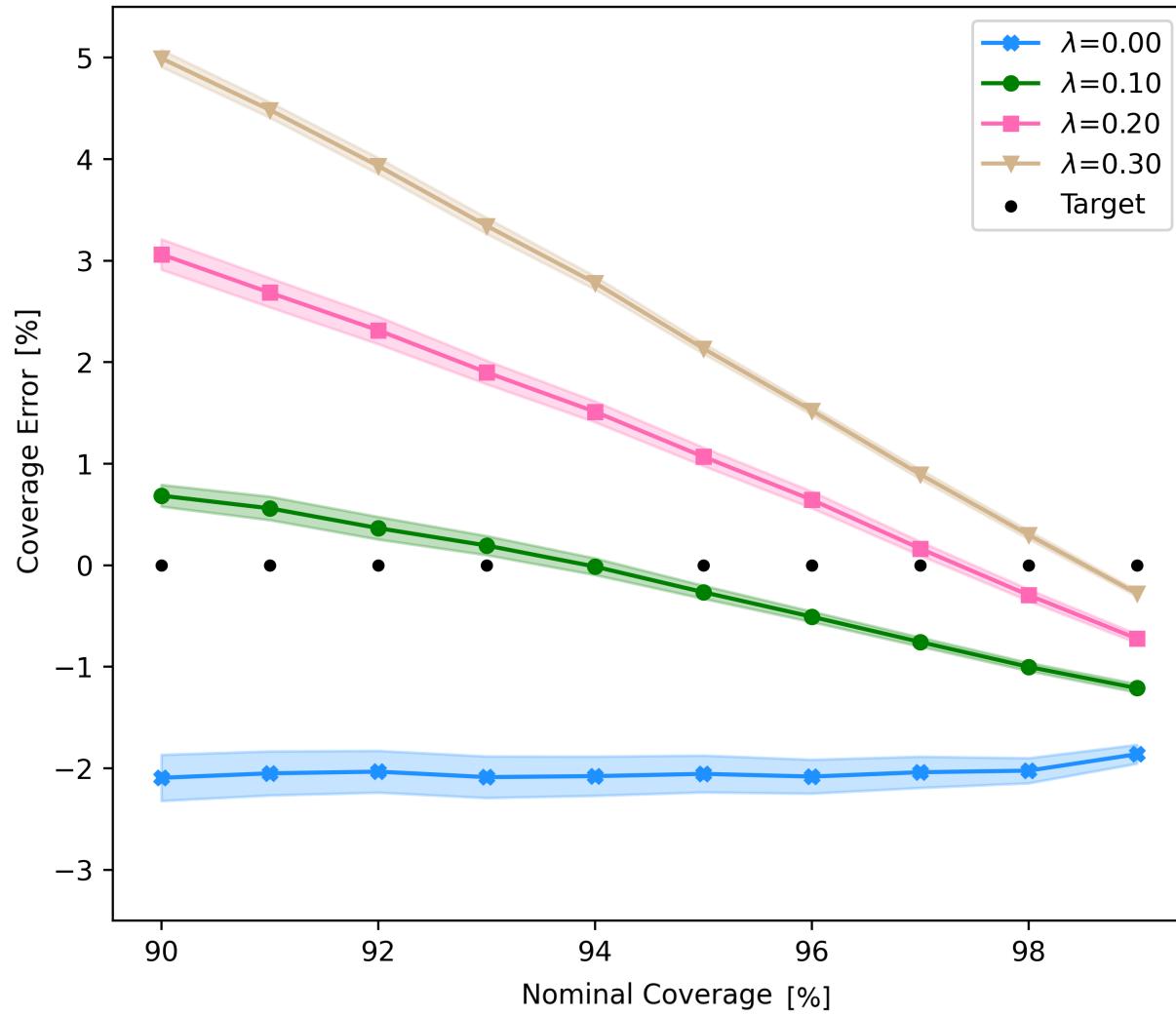


Figure: Prediction Intervals coverage plots with SE bands (10 repetitions)

# Numerical Results (on test set, 2012)

$\mathcal{G}$ -CRPS<sup>[ $\lambda$ ]</sup>

$\lambda$	MAPE [%]	RMSE [MWh]	CRPS [MWh]	EC(90%) [%]	EC(95%) [%]	AACE [%]
0.00	2.01 $\pm$ 0.01	409.58 $\pm$ 1.31	104.23 $\pm$ 0.33	87.90 $\pm$ 0.22	92.94 $\pm$ 0.18	2.03 $\pm$ 0.17
0.05	2.00 $\pm$ 0.01	408.45 $\pm$ 1.60	103.97 $\pm$ 0.39	89.21 $\pm$ 0.14	93.64 $\pm$ 0.17	1.26 $\pm$ 0.13
0.10	2.00 $\pm$ 0.01	406.49 $\pm$ 1.64	103.87 $\pm$ 0.43	90.68 $\pm$ 0.10	94.73 $\pm$ 0.06	0.59 $\pm$ 0.03
0.15	1.99 $\pm$ 0.01	406.11 $\pm$ 1.67	103.77 $\pm$ 0.42	91.87 $\pm$ 0.19	95.36 $\pm$ 0.11	0.91 $\pm$ 0.08
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$\mathcal{G}$ -MLE

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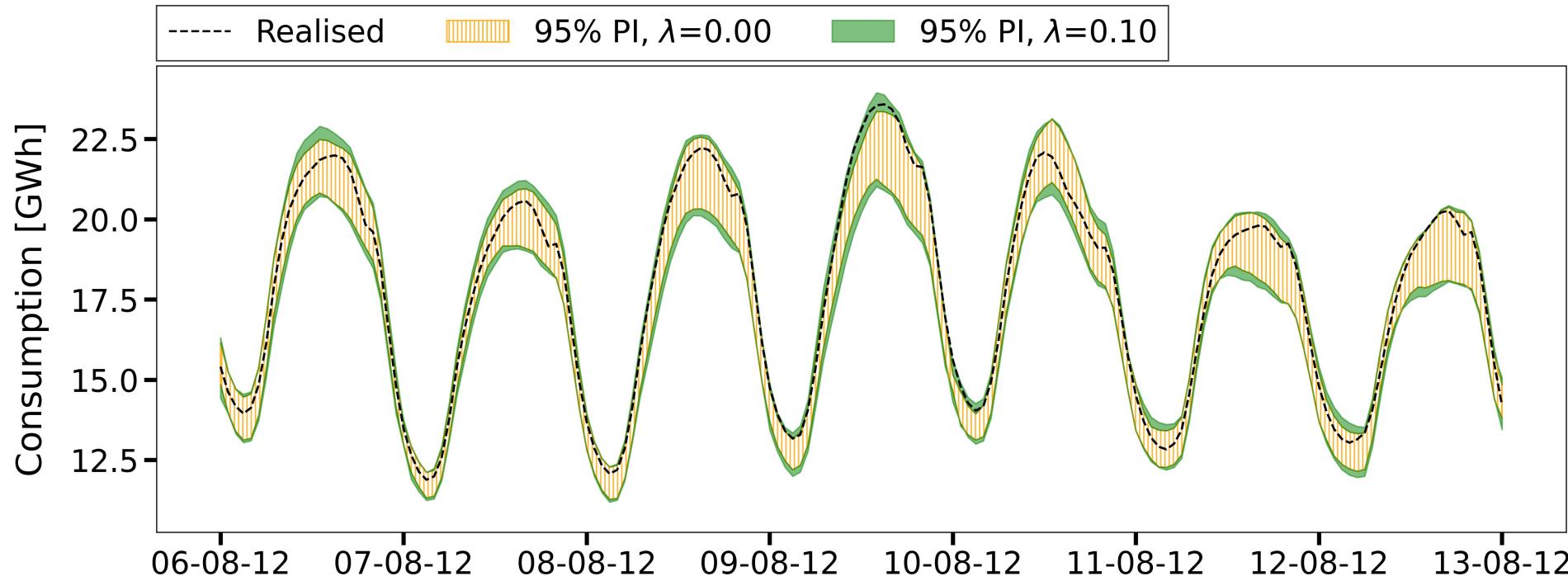


Figure: Modification of density forecasts of hourly demand on the test set (2012) for a week of August, when  $\mathcal{G}$ -CRPS<sup>[0.00]</sup>, the original CRPS loss function, and  $\mathcal{G}$ -CRPS<sup>[0.10]</sup>, the new proposed version of the CRPS, are used to train the predictive model.

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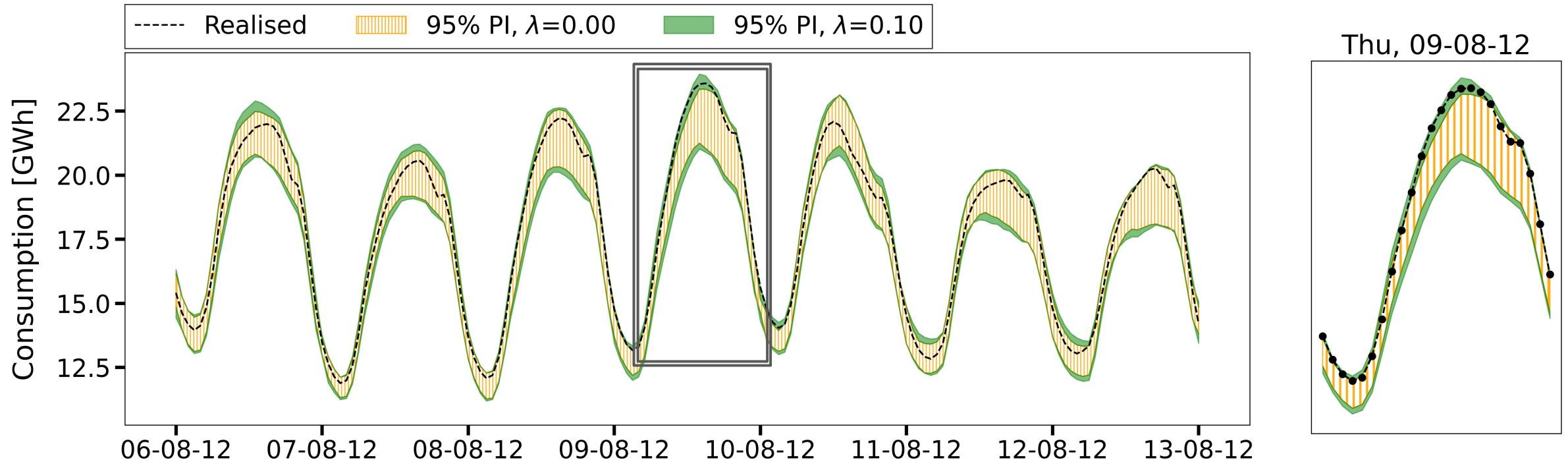


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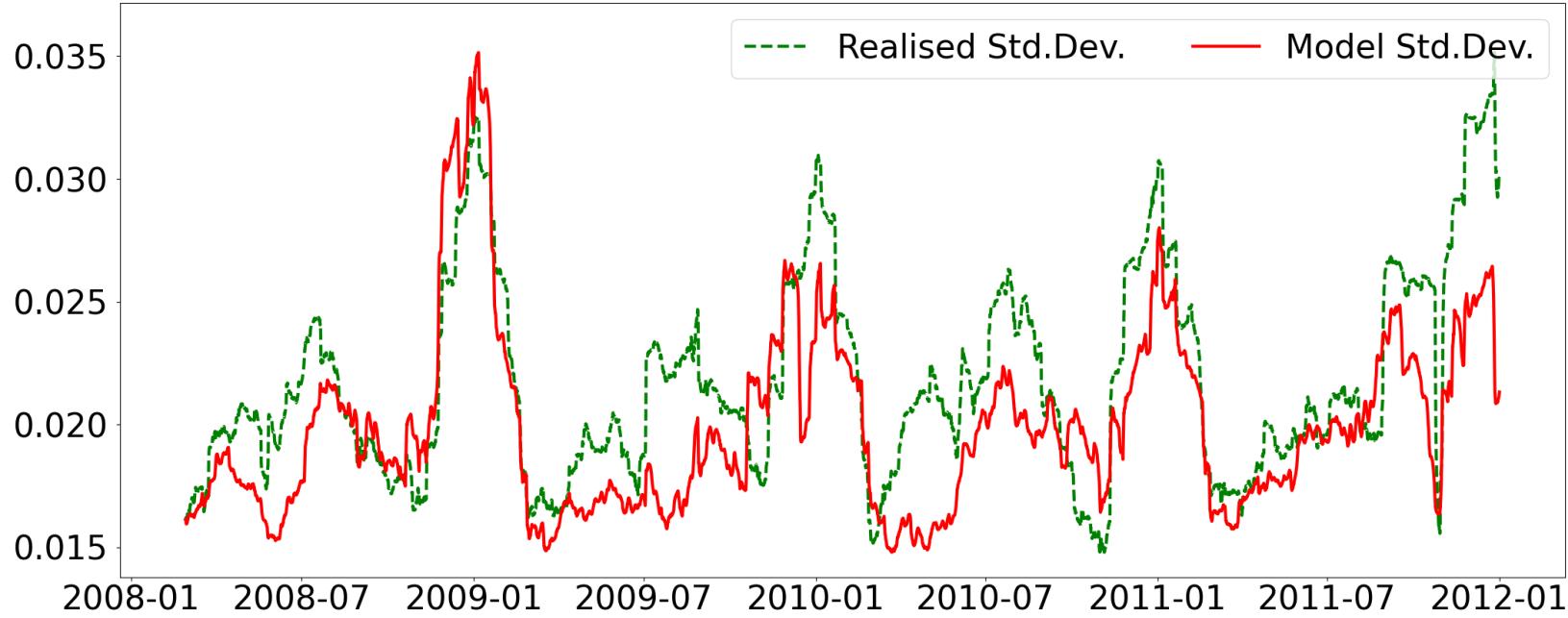
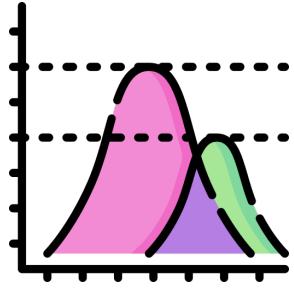


Figure: Realised standard deviation and standard deviation fitted by the RNN model when  $\mathcal{G}$ -CRPS<sup>[0.10]</sup> is used to train the predictive model. Standard deviation is computed considering a 2-months rolling-window approach.

# Conclusions

- i) We have designed a new parsimonious forecasting methodology for time series that display complex patterns (heteroskedasticity, autocorrelation, multi-seasonality)
- ii) We have designed a new family of loss functions to tackle predictive **overconfidence** in probabilistic time series forecasting
- iii) We have tested the methodology on a benchmark dataset. The predictive performance is **excellent** both in terms of **reliability** and point **accuracy**

# Essential bibliography

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