Tutorial on Conformal Predictive Distributions

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- **Conformal Predictive Distributions** (Vovk et al., 2017) are a novel approach to estimating the probability distribution of a continuous variable that depends on a number of features.
- CPDs probabilities correspond to long-term relative frequencies (within statistical fluctuation) under minimal assumptions.
 - CPDs require only that the data be generated independently by an unknown but fixed distribution.
 - No assumption on the type of distribution
 - No need of a prior
- Outline of the tutorial
 - Motivation, context
 - Predictive Distribution, Conformal Predictive Distribution
 - KRRPM
 - Evaluation of PD and real-life examples

The roots of prior-free predictive distributions



- Suppose data D is generated by a known distribution with unknown parameter θ , which we want to estimate.
- Bayesian statistics can provide a distribution for the parameter
 - Estimating a probability distribution comes naturally to Bayesian methods

 $p(heta | \mathcal{D}) \propto p(heta) p(\mathcal{D} | heta)$

- but you have to specify a prior $p(\theta)$.
- Frequentist statistics provides Confidence Intervals
 - Given $\alpha \in [0,1]$ coverage probability, we compute an interval

 $L(\mathcal{D}, \alpha), U(\mathcal{D}, \alpha)$

that contains the actual θ a fraction α of the time.

- $p(\theta|D)$ does not make sense in a strict frequentist framework.
- But frequentists recognized that predictive distributions would be useful!



- "The Holy Grail of parametric statistics" (Efron, 2010)
- Early attempts to arrive at prior-free posterior distributions can be traced back to Fisher's fiducial approach in the 1930s.
 - Not completely formalized; controversial; referred to as "Fisher's biggest blunder".
- There is now a resurgence of interest in the topic.
- Conformal Predictive Distributions are part of this trend.



- Unrestricted randomness, i.i.d. data
 - $\mathcal{D} = \{(x_1, y_1), \dots, (x_n, y_n)\}$ generated independently by fixed, but unknown P(X, Y)
- Intuitively, we seek an F(y, x) that has the properties of a Cumulative Distribution Function in y.

$$F(y, x, \mathcal{D}) = P\{Y \le y | X = x\}$$

As long as $(x_i, y_i) \sim P(X, Y)$, the intervals $(-\infty, y)$ contain y_i with relative frequency $F(y, x_i)$



• Let $F_X()$ be the (Cumulative) Distribution Function of the Random Variable X.

$$F_X(x) = P\{X \le x\}$$

• The RV that is obtained by evaluating the *F_X* on the RV *X* is uniformly distributed.

 $F_X(X) \sim U(0,1)$

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- Suppose we have observations z_1, \ldots, z_n from a set $\mathbf{Z} = \mathbf{X} \times \mathbb{R}$ and a test object $x_{n+1} \in \mathbf{X}$
- Let's call Predictive Distribution a function $Q(z_1, \ldots, z_n, (x_{n+1}, y))$ that
 - for any choice of training sequence z₁,..., z_n and any choice of test object x_{n+1} has the following properties of a CDF:
 - $Q(z_1, \ldots, z_n, (x_{n+1}, y))$ is monotonically increasing in y
 - $\lim_{y\to-\infty} Q(z_1,\ldots,z_n,(x_{n+1},y)) = 0$
 - $\lim_{y\to+\infty} Q(z_1,\ldots,z_n,(x_{n+1},y)) = 1.$
 - for any joint probability distribution P on Z,
 - $Q(z_1,...,z_n,z_{n+1}) \sim U$ when $(z_1,...,z_{n+1}) \sim P^{n+1}$.
- This definition omits some technicalities to keep things simple.

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- The "uniformity" property confers the ability to produce prediction intervals with **guaranteed coverage**
- Guaranteed coverage is the key property of predictive distributions
 - We can choose a confidence level α and we can read, off the predictive distribution, intervals of y in which the actual value falls with rate α (barring statistical fluctuation).



 Note that one can choose different prediction intervals for a given confidence level.



- One can choose the narrowest
 - i.e. where the slope of the predictive distribution is largest
- Or one around the median (previous slide)

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- The ECDF is guaranteed to converge to CDF for i.i.d. data (Glivenko-Cantelli theorem)
- The ECDF converges fast to the CDF (Dvoretzky–Kiefer–Wolfowitz inequality)
- To obtain a PDF you have to solve an ill-posed (unstable) problem:

$$\int h(x-t)f(t)dt = F(t) \qquad \text{where } h(x) \text{ is the step function}$$

• stable: a small variation in the right-side F(x) results in a small change in the solution f(x)



Probability predictions for 3 compounds



Pros

- Computing an ECDF from data is straightforward.
- You can read probabilities easily off the chart.

Cons

- Perceptually challenging to evaluate density
- You can't find the mode immediately.

Conformal Predictive Distributions

- It is possible to obtain Predictive Distributions by using a variant of Conformal Prediction for Regression.
- We define as conformity measure a function

$$A(z_1,\ldots,z_{n+1})$$

invariant w.r.t. permutations of the first *n* arguments.

• Given (x_{n+1}, y) , we compute conformity scores α_i^y as:

$$\alpha_i^{\mathcal{Y}} := A(z_1, \dots, z_{i-1}, z_{i+1}, \dots, z_n, (x_{n+1}, y), z_i), \qquad i = 1, \dots, n,$$

$$\alpha_{n+1}^{\mathcal{Y}} := A(z_1, \dots, z_n, (x_{n+1}, y)).$$

• Subject to some conditions on A(), the predictive distribution is then

$$Q(z_1,...,z_n,(x_{n+1},y)) := \frac{1}{n+1} \Big(\big| \{i=1,...,n+1 \mid \alpha_i^y < \alpha_{n+1}^y \} \big| \Big)$$



Requirements on the conformity measure



- Q() can be viewed as the estimate of probability under i.i.d. of drawing an observation with a smaller value of the CM than the hypothetical observation (x_{n+1}, y). It is as if we were testing the null hypothesis of i.i.d. using α as test statistic and computing the p-value Q().
- In contrast to CP, not all functions with the specified invariance property are conformity measures that result in valid predictive distributions.
- For the resulting Q() to have the properties of CDF, the conformity measure must be such that:
 - $\alpha_{n+1}^{y} \alpha_{i}^{y}$ is a monotonically increasing function of $y \in \mathbb{R}$

•
$$\lim_{y \to \pm \infty} (\alpha_{n+1}^y - \alpha_i^y) = \pm \infty$$

• A simple example is the $y - \hat{y}_{n+1}$ where \hat{y}_{n+1} is the estimate obtained with K nearest neighbours regression.

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Some details



- The definition was simplified for the sake of clarity.
- The properly defined CPD has a randomness element

$$Q(z_1,...,z_n,(x_{n+1},y)) := \frac{1}{n+1} \left(\left| \left\{ i = 1,...,n+1 \mid \alpha_i^y < \alpha_{n+1}^y \right\} \right| \right) + \frac{\tau}{n+1} \left(\left| \left\{ i = 1,...,n+1 \mid \alpha_i^y = \alpha_{n+1}^y \right\} \right| \right) \right.$$

where $au \sim U(0,1)$

• Informally, it's a "thick" distribution function, but the thickness is $\frac{1}{n+1}$, so it matters little as soon as you have a reasonably sized training set





• Let's use Kernel Ridge Regression

$$\hat{y}_{n+1} := k'(K + aI)^{-1}Y$$

where $k_i = \mathcal{K}(x_i, x_{n+1})$, $K_{i,j} := \mathcal{K}(x_i, x_j)$, i, j = 1, ..., n. Unfortunately, $y - \hat{y}_{n+1}$ is not a proper conformity measure. It fails to produce a strictly increasing function in y for high-leverage objects.

• Another possibility is to include the test object with the hypothetical label in the training set of the KRR.

$$\hat{\bar{y}}_{n+1} := \bar{k}'(\bar{K} + aI)^{-1}\bar{Y}$$

where $\bar{k}_i = \mathcal{K}(x_i, x_{n+1})$, $\bar{K}_{i,j} := \mathcal{K}(x_i, x_j)$, $\bar{y}_i = y_i$, i, j = 1, ..., n + 1. This too fails to guarantee a strictly increasing function in y.



• A conformity measure is obtained as:

$$\frac{y_{n+1} - \hat{\bar{y}}_{n+1}}{\sqrt{1 - \bar{h}_{n+1}}}$$

where \bar{h}_{n+1} is the element $\bar{H}_{n+1,n+1}$ of the hat matrix $\bar{H} := (\bar{K} + aI)^{-1}\bar{K}$

• There exists an explicit form of the resulting KRRPM. It can be implemented in a way that avoids recomputing from scratch the hat matrix for every test object.

Application of KRRPM to Chemoinformatics

- Data set from a study on an enzyme
 - 1368 compounds
 - 68 features (PhysChem properties)
- Training set: 1000 observations, randomly sampled
- KRRPM using Laplace kernel



• Validity

The predicted probabilities correspond to the long-term relative frequencies

• **Specificity** (a.k.a. Sharpness) The intervals for a given probability are as narrow as possible

- You can always make a valid predictive distribution: just output for all test objects the same PD, the ECDF of the label. But this would be a terrible forecast as it would have no specificity.
- CPDs guarantee validity under i.i.d. One can concentrate on improving specificity





• For probabilistic predictions, there are some established "scores"

• Brier loss:
$$\frac{1}{n} \sum_{i=1}^{n} (p_i - o_i)^2$$

- Log loss: $\frac{1}{n} \sum_{i=1}^{n} (o_i \log p_i + (1 o_i) \log(1 p_i))$
- These apply to tasks in which predictions must assign probabilities to a set of mutually exclusive **discrete** outcomes.
- They can be used on predictive distributions but they do not evaluate the PDs in their entirety.
- There are metrics and diagnostic tools for PDs
 - PIT, CRPS
 - Validity plot, Interval boxplots



- Probability Integral Transform (PIT): evaluate $F_i()$ on the actual label y_i
- If the predictions $F_i(y)$ are ideal, $F_i(y_i)$ are variates from a U(0, 1) distribution.
- The PIT can be used to check validity.
- The histogram of the PIT should be as flat as possible.
- Perhaps better, the ECDF of the PIT should be as close as possible to the (0,0)-(1,1) diagonal
 - Kolmogorov-Smirnov statistic and K-S test

Continuous Ranked Probability Score (CRPS)



• The quadratic measure of discrepancy between the forecast CDF F(y,x) and the "ideal forecast CDF" given the scalar observation y

$$\mathsf{CRPS}(F, x, y) = \int_{\mathbb{R}} \left[F(t, x) - \mathbb{I}(t \ge y)\right]^2 dt$$

where $\mathbb{I}()$ is the indicator function.



• For a number of predictions, one takes the average:

$$\overline{\mathsf{CRPS}}(F) = \frac{1}{n} \sum_{i=1}^{n} \mathsf{CRPS}(F, x_i, y_i)$$

Practical diagnostic tools: Validity plot



- Validity plot: actual coverage vs. confidence
- For all the confidence values of interest (e.g. $0.1, 0.2, \ldots, 0.9$)
 - compute the intervals for the objects in the validation set
 - compute the relative frequency of "interval contains actual label"
- The relative frequency should be close to the confidence



Practical diagnostic tools: Interval boxplots



- Descriptive statistics of the intervals
- The narrower the intervals, the more useful the predictions.



hERG Predicted Intervals (split random)

A real-life comparison



HLM Validity (split 3)





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- The validity guarantee rests on the i.i.d. assumption.
 - $\{(x_1, y_1), \dots, (x_n, y_n)\} \sim P^n(X, Y)$, where $(x_i, y_i) \sim P(X, Y)$
 - It is a minimal assumption of regularity
 - But it can be violated very easily in practice!

•
$$P(X, Y) = P(X) P(Y|X)$$

- P(X) varies: covariate shift
- P(Y|X) varies: in high-dimensional spaces, one visits just a small portion
- You should not assess validity and efficiency separately.
 - There is a trade-off between the two



- Conformal Predictive Distributions offer a prior-free, non-parametric way to estimate the distribution of random variable Y for an object x, based on previous data $(x_1, y_1), \ldots, (x_n, y_n)$.
- CPDs have a validity guarantee under a minimal assumption of test and training data being i.i.d.
 - CPD require only that the data be generated independently by an unknown but fixed distribution.
- KRRPM uses the flexible and regularized method of Kernel Ridge Regression to generate Conformal Predictive Distributions.



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 $\label{eq:prediction} \ensuremath{\mathsf{Prediction}}\xspace \ensuremath{\mathsf{with}}\xspace \ensuremath{\mathsf{Confidence-A}}\xspace \ensuremath{\mathsf{general}}\xspace \ensuremath{\mathsf{ramework}}\xspace \ensuremath{\mathsf{for}}\xspace \ensuremath{\mathsf{predictive}}\xspace \ensuremath{\mathsf{ramework}}\xspace \ensurem$

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